

SOME REMARKS ON THE LARGE SIEVE OF YU. V. LINNIK

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(Received November 22, 1967)

§ 1. Introduction

YU. V. LINNIK has discovered (see [1]) in 1941 a very powerful new method of elementary number theory, which he called the large sieve¹. In his original formulation the large sieve asserts that if we take any sequence S_N consisting of Z positive integers $\leq N$, and if Y denotes the number of those primes² $p \leq \sqrt{N}$ for which all the elements of the sequence S_N are contained in $\leq p(1 - \varepsilon)$ residue classes mod p , where $0 < \varepsilon < 1$, then one has

$$(1.1) \quad Y \leq \frac{20\pi N}{\varepsilon^2 Z}.$$

As shown by the second named author in [3], Linnik's method is capable to prove much more, namely that if Z is not too small compared with N , then the elements of the sequence S_N not only occupy "almost all" residue classes mod p with respect to most primes $p \leq \sqrt{N}$, but are almost uniformly distributed in the p residue classes mod p for most primes $p \leq \sqrt{N}$. More exactly, let us denote by $Z(a, p)$ (where $a = 0, 1, \dots, p-1$) the number of elements of the sequence S_N which are congruent to a mod p . Then one has, putting

$$(1.2) \quad \Delta^2(p) = p \sum_{a=0}^{p-1} \left(Z(a, p) - \frac{Z}{p} \right)^2.$$

the inequality³

¹ As regards important applications of the large sieve in number theory, see e.g. [2] 3], [4], [5], [6]; [5] and [6] contain many further references.

² In this paper p always denotes a prime number.

³ Here and in what follows all the constants of the O -estimates are absolute, i.e. do not depend on N , nor on the sequence S_N nor on Q .

$$(1.3) \quad \sum_{p \leq Q} \Delta^2(p) = O(Z^{2/3} N^{4/3} Q^{1/3})$$

for $Q \leq N^{3/5}$. Later, the second named author has found (see [7]), a new probabilistic method for proving theorems of the type of the large sieve. This method (developed further and generalized in the papers [8], [9], [10], [11], [12]) gave the result

$$(1.4) \quad \sum_{p \leq Q} \Delta^2(p) = O(Z(Q^3 + N))$$

for $Q \leq \sqrt{N}$. This estimate is better than (1.3) for $Q \leq N^{3/8}$, but weaker if $N^{3/8} < Q \leq \sqrt{N}$.

Especially for $Q = N^{1/3}$ this result gives

$$(1.5) \quad \sum_{p \leq N^{1/3}} \Delta^2(p) = O(NZ).$$

The estimate (1.5) is essentially best possible, because if for instance S_N is the sequence of odd numbers $\leq N$, one has $Z(0, 2) = 0$ and thus $2 \left(Z(0, 2) - \frac{Z}{2} \right)^2 = \frac{Z^2}{2}$ i.e. this single term is already of order NZ .

The probabilistic approach, besides leading to a very sharp estimate for $Q \leq N^{1/3}$, has thrown light on the reasons why an arbitrary sufficiently dense subsequence of the sequence $1, 2, \dots, N$ has to be almost uniformly distributed among the residue classes mod p for most $p \leq N^{1/3}$; it became obvious that this is due to the statistical independence (more exactly: almost independence) of the distribution mod p and mod q of the numbers $n \leq N$ for any two primes $p, q \leq N^{1/3}$, $p \neq q$.

In the last two years important progress was made on the large sieve. The first essential improvement was obtained by K. F. ROTH [13]. His result was sharpened by BOMBIERI [14] who has shown that (1.5) holds also for $Q = \sqrt{N}$. More exactly Bombieri proved

$$(1.6) \quad \sum_{p \leq Q} \Delta^2(p) = O(Z(Q^2 + N)).$$

Clearly (1.6) is superior to both (1.3) and (1.4) for the full ranges $Q \leq N^{3/5}$ resp. $Q \leq \sqrt{N}$.

An important generalization of Bombieri's theorem has been obtained by H. DAVENPORT and H. HALBERSTAM [15]. To make this advance clear one has to notice that putting

$$(1.7) \quad S(x) = \sum_{n \in S_N} e^{2\pi i n x}$$

one has

$$(1.8) \quad \Delta^2(p) = \sum_{a=1}^{p-1} \left| S\left(\frac{a}{p}\right) \right|^2.$$

Now DAVENPORT and HALBERSTAM have proved that if $\alpha_1, \alpha_2, \dots, \alpha_D$ are arbitrary real numbers in the interval $(0, 1)$ such that $|\alpha_i - \alpha_j| \cong \delta > 0$ for $i \neq j$, one has

$$(1.9) \quad \sum_{i=1}^D |S(\alpha_i)|^2 = O\left(Z\left(\frac{1}{\delta} + N\right)\right).$$

Clearly, if the numbers $\frac{a}{p}$ ($a = 1, 2, \dots, p-1$; $p \leq Q$) are taken as the numbers $\alpha_1, \dots, \alpha_D$ ($D = \sum_{p \leq Q} (p-1)$) then $\delta \cong \frac{1}{Q^2}$ and thus (in view of (1.8)) (1.9) implies (1.6).

Note that from (1.9) one obtains even more than (1.6), namely that

$$\sum_{q \leq Q} \sum_{\substack{a=1 \\ (a, q)=1}}^q \left| S\left(\frac{a}{q}\right) \right|^2 = O\left(Z\left(\frac{1}{\delta} + N\right)\right)$$

because if $(a, q) = 1$, $(a', q') = 1$ (here (a, q) denotes the greatest common divisor of a and q) one has for $q, q' \leq Q$ and $\frac{a}{q} \neq \frac{a'}{q'}$

$$\left| \frac{a}{q} - \frac{a'}{q'} \right| \cong \frac{1}{Q^2}.$$

Recently P. X. GALLAGHER [16] has found a very elegant and simple method for proving (1.9). More exactly, he proved

$$(1.10) \quad \sum_{\nu=1}^D |S(\alpha_\nu)|^2 \cong Z\left(\frac{1}{\delta} + \pi N\right)$$

which implies by (1.8)

$$(1.11) \quad \sum_{p \leq Q} A^2(p) \cong Z(Q^2 + \pi N).$$

Thus we have for $Q = \sqrt{N}$

$$(1.12) \quad \sum_{p \leq Q} A^2(p) \cong (\pi + 1)ZN.$$

In the paper [17] of the first named author it has been mentioned (without giving the proof in detail) that by a probabilistic argument it can be shown that (1.12) cannot hold if Q is of larger order of magnitude than $\sqrt{N \log N}$. The aim of the present paper is to prove this statement in detail, and to get some related results concerning the behaviour of $\sum_{p \leq Q} A^2(p)$, when S_N is a random subset of the set $\{1, 2, \dots, N\}$.

The results obtained throw some light on certain open problems connected with the large sieve.

§ 2. Equidistribution of random sequences in arithmetic progressions

In this § let S_N denote a random subsequence of the sequence $\{1, 2, \dots, N\}$ obtained as follows: let $\varepsilon_1, \varepsilon_2, \dots, \varepsilon_N$ be independent random variables, each of which takes on the values 1 and 0 with probability $\frac{1}{2}$; let S_N denote the set of those $n \leq N$ for which $\varepsilon_n = 1$. (It is easy to see that under these suppositions each of the 2^N subsets of the set $\{1, 2, \dots, N\}$ has the same probability to be chosen.) In this case

$$(2.1) \quad Z = \sum_{n=1}^N \varepsilon_n$$

$$(2.2) \quad Z(a, p) = \sum_{k=0}^{\left[\frac{N-a}{p}\right]} \varepsilon_{kp+a}$$

and consequently

$$(2.3) \quad \Delta^2(p) = p \cdot \sum_{a=0}^{p-1} \left(Z(a, p) - \frac{Z}{p} \right)^2$$

are all random variables. One obtains easily

$$(2.4) \quad \Delta^2(p) = p \cdot \sum_{a=0}^{p-1} Z^2(a, p) - Z^2$$

and thus, putting⁴

$$(2.5) \quad \pi_r(Q) = \sum_{p \leq Q} p^r \quad (r=0, 1, 2, \dots)$$

we have

$$(2.6) \quad R(Q) = \sum_{p \leq Q} \Delta^2(p) = \sum_{n=1}^N \sum_{m=1}^N [A_Q(n-m) - \pi_0(Q)] \varepsilon_n \varepsilon_m,$$

where

$$(2.7) \quad A_Q(k) = \sum_{\substack{p/k \\ p \leq Q}} p$$

and thus $A_Q(-k) = A_Q(k)$ and especially

$$(2.8) \quad A_Q(0) = \pi_1(Q).$$

Let us determine first the expectation⁵ of $R(Q)$. As

$$(2.9) \quad E(\varepsilon_n) = \frac{1}{2} \quad \text{and} \quad E(\varepsilon_n \varepsilon_m) = \begin{cases} \frac{1}{4} & \text{if } n \neq m \\ \frac{1}{2} & \text{if } n = m \end{cases}$$

⁴ Thus $\pi_0(Q)$ denotes the number of primes $\leq Q$.

⁵ The expectation of a random variable ξ will be denoted by $E(\xi)$.

we obtain

$$(2.10) \quad E(R(Q)) = \frac{1}{4} \sum_{n=1}^N \sum_{m=1}^N (A_Q(n-m) - \pi_0(Q)) + \frac{N}{4} \pi_1(Q).$$

Now clearly

$$(2.11) \quad \sum_{n=1}^N \sum_{m=1}^N [A_Q(n-m) - \pi_0(Q)] = \sum_{p \leq Q} p \left\{ \sum_{a=0}^{p-1} \left[\left[\frac{N-a}{p} \right] + 1 \right]^2 - \frac{N^2}{p^2} \right\}.$$

Let us suppose that $N \equiv r \pmod{p}$, where $0 \leq r < p$. Then we have

$$(2.12) \quad p \cdot \left\{ \sum_{a=0}^{p-1} \left[\left[\frac{N-a}{p} \right] + 1 \right]^2 \right\} = 2N + r(p-r) + p - 2r.$$

Thus it follows that

$$(2.13) \quad \sum_{n=1}^N \sum_{m=1}^N [A_Q(n-m) - \pi_0(Q)] = 2N\pi_0(Q) + O(Q^2\pi_0(Q))$$

and thus, taking into account that $\pi_0(Q) = \frac{Q}{\log Q} + O\left(\frac{Q}{\log^2 Q}\right)$ and

$$(2.14) \quad \pi_1(Q) = \frac{Q^2}{2 \log Q} + O\left(\frac{Q^2}{\log^2 Q}\right)$$

it follows

$$(2.15) \quad E(R(Q)) = \frac{NQ\pi_0(Q)}{8} + O(Q^2\pi_0(Q)) + O\left(\frac{Q^2N}{\log^2 Q}\right).$$

Thus the expectation of $R(Q)$ is smaller by a factor of order $\frac{1}{\log Q}$ as NQ^2 .

Note that the expectation of $R(Q)$ can be interpreted as its average over all 2^N subsequences of the sequence $\{1, 2, \dots, N\}$. Thus the average of $R(Q)$ is of order $O(N^2)$ even for $Q = O(\sqrt{N \log N})$ while for its maximum according to (1.6) this is known only for $Q = O(\sqrt{N})$. It is an open question whether the estimate

$$(2.16) \quad R(Q) = O(N^2)$$

holds for all sequences S_N if $Q \sim \sqrt{N}\psi(N)$ for some function $\psi(N)$ such that $\psi(N) \rightarrow \infty$ for $N \rightarrow \infty$. Our method is not capable of giving such a result; however by evaluating the variance of the random variable $R(Q)$ we can show by Čebishev's inequality that the estimate (2.16) is valid at least for most subsequences S_N .

To evaluate the variance⁶ of $R(Q)$ note that though the random variables $\varepsilon_n \varepsilon_m$ are not independent, they are pairwise uncorrelated and thus the variance

⁶ The variance of a random variable ξ will be denoted by $D^2(\xi)$.

of the sum on the right hand side of (2.6) is equal to the sum of the variances of the single terms. As $D^2(\varepsilon_n \varepsilon_m) = \frac{3}{16}$ if $n \neq m$ and $D^2(\varepsilon_n^2) = \frac{4}{16}$.

$$(2.17) \quad D^2(R(Q)) = \frac{3}{16} \sum_{n=1}^N \sum_{m=1}^N (A_Q(n-m) - \pi_0(Q))^2 + \frac{N}{16} (\pi_1(Q) - \pi_0(Q))^2.$$

Now clearly

$$(2.18) \quad \sum_{n=1}^N \sum_{m=1}^N A_Q^2(n-m) \cong N^2[\pi_0^2(Q) + \pi_1(Q) - \pi_0(Q)] + 2N\pi_1^2(Q) + \pi_2^2(Q).$$

As further from (2.11) we have

$$\sum_{n=1}^N \sum_{m=1}^N A_Q(n-m) \cong N^2\pi_0(Q) - 2N\pi_0(Q) + \pi_1(Q)$$

it follows

$$(2.19) \quad D^2(R(Q)) \cong \frac{3}{16} N^2(\pi_1(Q) - \pi_0(Q)) + \\ + \frac{3N}{8} \left(\pi_1^2(Q) + 2\pi_0^2(Q) + \frac{(\pi_1(Q) - \pi_0(Q))^2}{6} \right) + \frac{3}{16} \pi_2^2(Q) - \frac{3}{8} \pi_1(Q)\pi_0(Q).$$

In view of (2.14), it follows

$$(2.20) \quad D^2(R(Q)) = O\left(\frac{N^2Q^2}{\log Q}\right) + O\left(\frac{NQ^4}{\log^2 Q}\right)$$

i.e.

$$(2.21) \quad \frac{D(R(Q))}{E(R(Q))} = O\left(\frac{\sqrt{\log Q}}{Q}\right) + O\left(\frac{1}{\sqrt{N}}\right).$$

It follows from Čebishev's inequality that for $\lambda > 1$ with probability $\cong 1 - \frac{1}{\lambda^2}$ $R(Q)$ is contained in an interval

$$[E(R(Q)) - \lambda D(R(Q)), E(R(Q)) + \lambda D(R(Q))].$$

Choosing for λ the value $\lambda = \min\left(\frac{Q}{(\log Q)^{3/2}}, \frac{\sqrt{N}}{\log Q}\right)$ it follows that for all but $\frac{2^N}{\lambda^2}$ possible exceptions for all other sequences, i.e. for the large majority of all sequences, $R(Q)$ is of order $\frac{NQ^2}{8 \log Q} + O\left(\frac{NQ^2}{\log^2 Q}\right)$.

Thus we have proved the following

THEOREM 1. Let us consider all 2^N subsequences S_N of the sequence $\{1, 2, \dots, \dots, N\}$. We have for all these subsequences with the possible exception of $\frac{2^N}{\lambda^2}$ such sequences

$$(2.22) \quad R(Q) = \frac{NQ^2}{8 \log Q} + O\left(\frac{NQ^2}{\log^2 Q}\right)$$

where $Q \cong N^{1/3}$ and

$$(2.23) \quad \lambda = \min\left(\frac{Q}{(\log Q)^{3/2}}, \frac{\sqrt{N}}{\log Q}\right).$$

Thus, if $Q \cong \sqrt{N \log N}$, (2.22) holds except for at most $\frac{2^N \log^3 Q}{Q^2}$ sequences, while for $Q > \sqrt{N \log N}$ (2.22) holds, except for at most $\frac{2^N \log^2 N}{N}$ sequences.

COROLLARY. If $Q = \sqrt{AN \log N}$ ($A > 1$) then $R(Q) \sim \frac{AN^2}{8}$ except for at most $\frac{2^N \log^2 N}{N}$ sequences.

Let us now consider the quantity

$$\text{Max}_{p \cong Q} \left[\text{Max}_{0 \cong a \cong p-1} \left| Z(a, p) - \frac{Z}{p} \right| \right].$$

It is easy to show [using the central limit theorem and the fact that for any given p the quantities $Z(a, p)$ ($a = 0, 1, \dots, p-1$) are independent], that

$$(2.24) \quad P\left(\text{Max}_{0 \cong a \cong p-1} \left| Z(a, p) - \frac{Z}{p} \right| > \sqrt{\frac{N \log pQ}{2p}}\right) = O\left(\frac{1}{Q}\right)$$

and thus except for at most $O\left(\frac{2^N}{\log Q}\right)$ exceptional sequences we have

$$\left| Z(a, p) - \frac{Z}{p} \right| \cong \sqrt{\frac{N \log pQ}{2p}}$$

for all a and p ($0 \cong a \cong p-1, p \cong Q$).

On the other hand, using again the independence of the random variables $Z(a, p)$ ($a = 0, 1, \dots, p-1$) and the central limit theorem it follows that for all except for at most $O\left(\frac{2^N}{N}\right)$ sequences S^N , one has, for all p such that

$$C \log N < p < \frac{N}{\sqrt{\log N}}$$

$$\Delta^2(p) = \frac{Np}{4} \left(1 + O\left(\frac{1}{\sqrt{\log N}}\right)\right)$$

if C is a sufficiently large positive number.

§ 3. The values of a random trigonometrical polynomial at well spaced points

In this § we shall consider the sum

$$(3.1) \quad T(\alpha, S) = \sum_{\nu=1}^D |S(\alpha_\nu)|^2$$

where $\alpha_1, \alpha_2, \dots, \alpha_D$ are real numbers "well spaced" in the sense of Davenport and Halberstam, satisfying

$$(3.2) \quad 0 < \alpha_1 < \alpha_2 < \dots < \alpha_D < 1 \quad \text{and} \\ \alpha_{\nu+1} - \alpha_\nu \geq \delta > 0 \quad \text{for } \nu = 1, 2, \dots, D-1$$

and $S(\alpha)$ is the random trigonometric polynomial

$$(3.3) \quad S(\alpha) = \sum_{n=1}^N \varepsilon_n e^{2\pi i n \alpha}$$

where $\varepsilon_1, \dots, \varepsilon_N$ are independent random variables, each taking on the values 1 and 0 with probability $\frac{1}{2}$.

We first evaluate the expectation of $T(\alpha, S)$. We have clearly

$$(3.4) \quad T(\alpha, S) = \sum_{n=1}^N \sum_{m=1}^N \varepsilon_n \varepsilon_m \sum_{\nu=1}^D e^{2\pi i (n-m)\alpha_\nu}$$

and thus

$$(3.5) \quad E(T(\alpha, S)) = \frac{1}{2} \sum_{\nu=1}^D \left(\frac{N}{2} + \sum_{l=1}^{N-1} (N-l) \cos 2\pi l \alpha_\nu \right) + \frac{ND}{4}.$$

Now it is well known that

$$(3.6) \quad \frac{N}{2} + \sum_{l=1}^{N-1} (N-l) \cos 2\pi l \alpha_\nu = \frac{\sin^2 N\pi \alpha_\nu}{2 \sin^2 \pi \alpha_\nu}.$$

As a matter of fact the formula (3.6) is well known as a formula for Fejér's kernel of the arithmetic means of Fourier series.

It follows from (3.5) and (3.6) that

$$(3.7) \quad E(T(\alpha, S)) = \frac{1}{4} \sum_{\nu=1}^D \frac{\sin^2 N\pi \alpha_\nu}{\sin^2 \pi \alpha_\nu} + \frac{ND}{4}.$$

Let us now consider the special case when $\alpha_Q^* = (\alpha_1^*, \dots, \alpha_D^*)$ is the set of all numbers $\frac{a}{q}$ with $(a, q) = 1$, $1 \leq a \leq q$, $1 < q \leq Q \leq N$. It is easy to see that

$$(3.8) \quad \sum_{q \leq Q} \sum_{\substack{a=1 \\ (a,q)=1}}^q \frac{\sin^2 N\pi \frac{a}{q}}{\sin^2 \pi \frac{a}{q}} = O(Q^3)$$

thus, denoting by $\varphi(q)$ the number of numbers $a < q$ relatively prime to q , we have

$$(3.9) \quad E(T(\alpha_Q^*, S)) = \frac{N}{4} \sum_{q=1}^Q \varphi(q) + O(Q^3).$$

As however

$$(3.10) \quad \sum_{q=1}^Q \varphi(q) = \frac{3Q^2}{\pi^2} + O(Q \log Q)$$

it follows that

$$(3.11) \quad E(T(\alpha_Q^*, S)) = \frac{3Q^2N}{4\pi^2} + O(NQ \log Q) + O(Q^3).$$

It follows that for $Q = o(N)$ there exists for each $\varepsilon > 0$ a sequence S_N for which

$$(3.12) \quad T(\alpha_Q^*, S_N) > \frac{3Q^2N(1-\varepsilon)}{4\pi^2}.$$

Thus the estimate

$$(3.13) \quad \sum_{q \leq Q} \sum_{\substack{a=1 \\ (a,q)=1}}^q \left| S\left(\frac{a}{q}\right) \right|^2 = O(N^2)$$

which according to the theorem of Davenport and Halberstam is valid for $Q \leq \sqrt{N}$ cannot be valid if Q is of larger order of magnitude than \sqrt{N} . By evaluating the variance of $T(\alpha, S)$ one can prove even more, namely that $T(\alpha_Q^*, S_N) \sim \frac{3Q^2N}{4\pi^2}$ for all except $o(2^N)$ sequences S_N , if $\frac{1}{Q} = o(1)$ and $\frac{Q}{N} = o(1)$. In particular one can prove that

$$(3.14) \quad D \left(\sum_{q \leq \sqrt{N}} \sum_{\substack{a=1 \\ (a,q)=1}}^{q-1} \left| S\left(\frac{a}{q}\right) \right|^2 \right) = O(N^{3/2})$$

which implies that except for at most $O\left(\frac{2^N \log N}{N}\right)$ exceptional sequences

$$(3.15) \quad \sum_{q \leq \sqrt{N}} \sum_{\substack{a=1 \\ (a,q)=1}}^{q-1} \left| S\left(\frac{a}{q}\right) \right|^2 \sim \frac{3N^2}{4\pi^2}.$$

Let us summarize now our results: Theorem 1 shows that the estimate

$$(3.16) \quad R(Q) = O(N^2)$$

cannot hold if Q is of larger order of magnitude than $\sqrt{N \log N}$. It remains an open question whether (3.16) holds if $\sqrt{N} \leq Q \leq \sqrt{N \log N}$. However, (3.11) shows that even if (3.16) is true for the range $\sqrt{N} \leq Q \leq \sqrt{N \log N}$ it cannot be proved by the methods used up to now, as all these methods gave estimates for $R(Q)$ through estimating $T(\alpha_Q^*, S)$.

§ 4. Some open problems

Let S_N denote a subsequence of the sequence $\{1, 2, \dots, N\}$ which contains at least cN elements ($0 < c < 1$). Let $Y(\alpha, \varepsilon)$ where $0 < \varepsilon < 1$ and $1/2 \leq \alpha < 1$ denote the number of those primes $p \leq N^\alpha$ for which at least $p\varepsilon$ residue classes mod p do not contain any element of S_N . It follows already from Linnik's result (1.1) that $Y\left(\frac{1}{2}, \varepsilon\right)$ is bounded, namely that

$$(4.1) \quad Y\left(\frac{1}{2}, \varepsilon\right) \leq \frac{20\pi}{c\varepsilon^2}.$$

From (1.12) one obtains the slightly better estimate

$$(4.2) \quad Y\left(\frac{1}{2}, \varepsilon\right) \leq \frac{\pi + 1}{\varepsilon c}.$$

As regards $Y(\alpha, \varepsilon)$ with $1/2 < \alpha < 1$ we get from (1.11) the estimate

$$(4.3) \quad Y(\alpha, \varepsilon) \leq \frac{N^{2\alpha-1}}{\varepsilon c} + \frac{\pi}{\varepsilon c}.$$

It seems probable that (4.3) is far from being best possible; it is an open problem whether $Y(\alpha, \varepsilon)$ is bounded for every α with $\frac{1}{2} < \alpha < 1$, or not. Of course, $Y(1, \varepsilon)$ is not bounded: as a matter of fact if S_N is the sequence of numbers $\leq Nc$ ($0 < c < \frac{1}{2}$) and $0 < \varepsilon < \frac{1}{2}$ then for all primes p with $\frac{cN}{1-\varepsilon} < p < N$ at least $p\varepsilon$ residue classes mod p do not contain any element of S_N , and thus

$$Y(1, \varepsilon) \geq \frac{N}{\log N} \left(1 - \frac{c}{1-\varepsilon}\right) + O\left(\frac{N}{\log^2 N}\right).$$

Another related problem is the following: if $0 < \varepsilon < 1$ let S_N be a subsequence of the sequence $\{1, 2, \dots, N\}$ such that for every p with $A_\varepsilon < p < N^\alpha$ where $A_\varepsilon > 0$, $0 < \alpha < 1$ there are at least εp residue classes mod p which do not contain any element of S_N . What is the maximum $M_N(\varepsilon, \alpha)$ of the number of terms of such a sequence S_N ? It is easy to show that for each ε with $0 < \varepsilon < 1/2$ $M_N(\varepsilon, 1) \geq [\sqrt{N}]$. As a matter of fact let S_N denote the sequence of squares $\leq N$. Clearly if b is a quadratic non-residue mod p , then there is no element of the sequence $1^2, 2^2, \dots, k^2, \dots$ which is congruent to b mod p ; thus for each p the number of empty residue classes is at least $\frac{p-1}{2}$ if $p \geq 3$.

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